

Lois usuelles - Espérance / Variance

Escoffier

I Loi uniforme.

$X \rightsquigarrow \mathcal{U}(\llbracket 1, n \rrbracket)$ si $X(\omega) = \llbracket 1, n \rrbracket$

$$\forall i \in \llbracket 1, n \rrbracket \quad \mathbb{P}(X=i) = \frac{1}{n}$$

$$E(X) = \frac{n+1}{2} \quad \text{et} \quad V(X) = \frac{n^2-1}{12}$$

preuve $E(X) = \sum_{i=1}^n i \times \frac{1}{n} = \frac{1}{n} \times \frac{n(n+1)}{2} = \frac{n+1}{2}$

$$E(X^2) = \sum_{i=1}^n i^2 \times \frac{1}{n} = \frac{1}{n} \frac{n(n+1)(2n+1)}{6} = \frac{(n+1)(2n+1)}{6}$$

$$\begin{aligned} V(X) &= E(X^2) - (E(X))^2 = \frac{(n+1)(2n+1)}{6} - \frac{(n+1)^2}{4} = \frac{(n+1)(2(2n+1) - 3(n+1))}{12} \\ &= \frac{(n+1)(4n-2)}{12} = \frac{(n+1)(n-1)}{12} \end{aligned}$$

II Bernoulli.

$$p \in]0, 1[, X \rightsquigarrow \mathcal{B}(p) \quad \begin{cases} \mathbb{P}(X=0) = 1-p \\ \mathbb{P}(X=1) = p \end{cases}$$

$$E(X) = p \quad V(X) = p(1-p)$$

preuve : $E(X) = 0 \times \mathbb{P}(X=0) + 1 \times \mathbb{P}(X=1) = p$

$$E(X^2) = 0 \times \mathbb{P}(X=0) + 1^2 \times \mathbb{P}(X=1) = p$$

$$V(X) = E(X^2) - (E(X))^2 = p - p^2 = p(1-p)$$

III Binomial.

$n \in \mathbb{N}^*$ $p \in]0, 1[$ $X \rightsquigarrow \mathcal{B}(n, p)$, $X(\omega) = \llbracket 0, n \rrbracket$

$$\forall k \in \llbracket 0, n \rrbracket \quad \mathbb{P}(X=k) = \binom{n}{k} p^k (1-p)^{n-k}$$

$$E(X) = np \text{ et } V(X) = np(1-p)$$

$$\begin{aligned} E(X) &= \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k} \\ &= \sum_{k=1}^n k \binom{n-1}{k-1} p^k (1-p)^{n-k} \\ &= np \sum_{k=1}^n \binom{n-1}{k-1} p^{k-1} (1-p)^{(n-1)-(k-1)} \quad (\text{Binôme}) \\ &= np \end{aligned}$$

$$\text{Pour la variance } k^2 = k(k-1) + k.$$

IV Loi de Poisson.

$\lambda > 0$, X na $\mathcal{P}(\lambda)$ si $X(\omega) = n$ et $\forall n \in \mathbb{N}$ $P(X=n) = \frac{\lambda^n}{n!} e^{-\lambda}$

\Rightarrow Loi limite d'une binomiale $B(n, \frac{\lambda}{n})$ avec $\lambda = np$

$$E(X) = \lambda \text{ et } V(X) = \lambda$$

$$E(X) = \sum_{k=0}^{+\infty} k e^{-\lambda} \frac{\lambda^k}{k!} = e^{-\lambda} \sum_{k=1}^{+\infty} \frac{\lambda^k}{(k-1)!} = e^{-\lambda} \lambda e^{\lambda} = \lambda \quad (\text{c'est par d'habitude})$$

$$\begin{aligned} E(X^2) &= \sum_{k=0}^{+\infty} k^2 e^{-\lambda} \frac{\lambda^k}{k!} = \sum_{k=0}^{+\infty} (k(k-1) + k) e^{-\lambda} \frac{\lambda^k}{k!} \\ &= \sum_{k=0}^{+\infty} k(k-1) e^{-\lambda} \frac{\lambda^k}{k!} + \sum_{k=0}^{+\infty} k e^{-\lambda} \frac{\lambda^k}{k!} \\ &= e^{-\lambda} \sum_{k=2}^{+\infty} \frac{\lambda^k}{(k-2)!} + \lambda \\ &= e^{-\lambda} \lambda^2 e^{\lambda} + \lambda = \lambda^2 + \lambda \end{aligned}$$

$$\text{et } V(X) = E(X^2) - E(X)^2 = \lambda.$$